Statistics

Chapter 4

Section 4-4 Comparison of Means with Student's *t*

t Test: Comparison of Means

Are the *means* of two sets of measurements "statistically different"?

If you make two sets of measurements of the same quantity, $\overline{x}_1 \neq \overline{x}_2$ generally ndom variations in measurements.

The *t* test determines if there is a statistical difference betwe \overline{x}_1 and \overline{x}_2 .

If $t_{\text{calculated}} > t_{\text{table}}$, then reject the null hypothesis.

- There is <5% chance that the two data sets came from populations with the same population mean.
- The difference is considered significant.

Comparisons of Means with Student's t

Case 1

Comparex to a known value:

- Measure quantity several times.
- Obtain \overline{x} and s.

Does \overline{x} compare to accepted answer, μ ?

Case 2

Compare \overline{x}_1 to \overline{x}_2 with replicate samples:

- Measure quantity multiple times by two different methods.
- Obtain $\overline{x}_1 \pm s_1$ and $\overline{x}_2 \pm s_2$ (for each method).

Does \overline{x}_1 agree with \overline{x}_2 within experimental uncertainty?

Case 3 (paired test)

Compare two methods where samples are not duplicated:

- Measure sample A once by method 1 and once by method 2.
- Measure sample B once by method 1 and once by method 2.

Do the two methods agree within experimental uncertainty?

Case 1: Comparing Measured Result with "Known" Value

A coal sample is certified to contain 3.19 wt% sulfur. A new analytical method measures values of 3.29, 3.22, 3.30, and 3.23 wt% sulfur, giving a measures ∂_0 and a standard deviation s = 0.041.

Does the answer using the new method agree with the known answer? Confidence Interval = $\overline{x} \pm \frac{ts}{\sqrt{n}} = 3.26_0 \pm \frac{(3.182)(0.004 \ 1)}{\sqrt{4}} = 3.260 \pm 0.006 \ 5$

- Confidence interval = 3.19_5 to 3.32_5 wt%
- Known value, **3.19 wt%**, outside 95% confidence interval
- Method gives "different" result from known result (Result is so close scientist might want to complete a few more trials to confirm)

Case 2a: Comparing Replicate Measurements When Standard Deviations Are *Not* Significantly Different

Recall HCO_3^- in horse blood is measured after each

: race.	Original instrument	Substitute instrument
Mean (\bar{x} , mM)	36.14	36.20
Standard deviation (s, mM)	0.28	0.47
Number of measurements (<i>n</i>)	10	4

$$F_{\text{calculated}} = \frac{s_1^2}{s_2^2} = \frac{(0.47)^2}{(0.28)^2} = 2.8_2$$

 $F_{calculated} = 2.8_2 < F_{table} = 5.08$, so s_1 and s_2 are **not** significantly different.

Do the means of the two methods

differ?

$$t = \frac{|\overline{x}_1 - \overline{x}_2|}{s_{\text{pooled}}} \sqrt{\frac{n_1 n_2}{n_1 + n_2}}$$

$$s_{\text{pooled}} = \sqrt{\frac{\sum_{1} (x_{i} - \overline{x}_{1})^{2} + \sum_{2} (x_{j} - \overline{x}_{2})^{2}}{n_{1} + n_{2} - 2}}$$
$$= \sqrt{\frac{s_{1}^{2} (n_{1} - 1) + s_{2}^{2} (n_{2} - 1)}{n_{1} + n_{2} - 2}}$$

Case 2a: Comparing Replicate Measurements When Standard Deviations Are *Not* Significantly Different

Do the means of the two methods differ?		Original instrument	Substitute instrument
2(1, 1) + 2(1, 1) $0.201(10, 1) + 0.072(1, 1)$	(<i>x</i> , mM)	36.14	36.20
$s_{\text{pooled}} = \sqrt{\frac{s_1(n_1-1)+s_2(n_2-1)}{10+1}} = \sqrt{\frac{0.28(10-1)+0.47(4-1)}{10+1}} = 0.33_8$	(s, mM)	0.28	0.47
$V n_1 + n_2 - 2 V 10 + 4 - 2$	(<i>n</i>)	10	4
$t = \frac{ \bar{x}_1 - \bar{x}_2 }{s_{\text{pooled}}} \sqrt{\frac{n_1 n_2}{n_1 + n_2}} = \frac{ 36.14 - 35.20 }{0.33_8} \sqrt{\frac{10 \times 4}{10 + 4}} = 0.30_0$	95% critie (n ₁ + n ₂ – lies b	cal value of <i>t</i> ir 2) = 12 degree etween 2.228 a	n Table 4-4 for es of freedom and 2.131

If $t_{\text{calculated}} > t_{\text{table}}$, then reject the null hypothesis.

 $t_{\rm calculated}(0.300) < t_{\rm table}(2.131)$

- There is <5% chance that the two data sets came from populations with the same population mean.</p>
- The difference in means is *not* considered significant.

Lord Rayleigh and the Discovery of Argon

Dry air is composed of ~1/5 oxygen and ~4/5 nitrogen

Measured nitrogen with two experiments:

(at constant temperature, pressure, and volume)

- Mass of N₂ after removing O₂ from air
- Mass of N₂ generated from chemical decomposition

Do the means of the two methods differ?



Table 4-5

From air (g)	From Chemical decomposition (g)
2.310 17	2.301 43
2.309 86	2.298 90
2.310 10	2.298 16
2.310 01	2.301 82
2.310 24	2.298 69
2.310 10	2.299 40
2.310 28	2.298 49
—	2.298 89
Average 2.310 10 ₉	2.299 47 ₂
Standard deviation	
0.000 14 ₃	0.001 37 ₉

Case 2b: Comparing Replicate Measurements When Standard Deviations *Are* Significantly Different

Do the means of the two methods differ?

If the standard deviations of the two methods differ (*F*-test), the *t* test equations become:

$$t_{\text{calculated}} = \frac{|\bar{x}_{1} - \bar{x}_{2}|}{\sqrt{(s_{1}^{2}/n_{1}) + (s_{2}^{2}/n_{2})}} = \frac{|\bar{x}_{1} - \bar{x}_{2}|}{\sqrt{(u_{1}^{2}) + (u_{2}^{2})}}$$
Standard deviation of the mean
$$u_{i} = s_{i}/\sqrt{n_{i}}$$
Degrees of freedom
$$= \frac{(s_{1}^{2}/n_{1} + s_{2}^{2}/n_{2})^{2}}{(s_{1}^{2}/n_{1})^{2}} + \frac{(s_{2}^{2}/n_{2})^{2}}{n_{2}-1} = \frac{(u_{1}^{2} + u_{2}^{2})^{2}}{(u_{1}^{2} - 1)^{2}}$$
Standard deviation of the mean
$$u_{i} = s_{i}/\sqrt{n_{i}}$$

Round degrees of freedom to the nearest integer.

Compare $t_{calculated}$ to t_{table} at 95% confidence using appropriate degrees of freedom.

Example: Is Rayleigh's N_2 from Air Denser Than N_2 from Chemicals? (1 of 4)

The average mass of nitrogen from air is $\overline{x}_1 = 2.310 \, 10_9$ g, with a standard deviation of $s_1 = 0.000 \, 14_3$ (for $n_1 = 7$ measurements). The average mass from chemical decomposition is $\overline{x}_2 = 2.299 \, 47_2$ g, with a standard deviation of $s_2 = 0.001 \, 37_9$ (for $n_2 = 8$ measurements). Are the two masses significantly different?

Table 4-5

From air (g)	From Chemical decomposition (g)
2.310 17	2.301 43
2.309 86	2.298 90
2.310 10	2.298 16
2.310 01	2.301 82
2.310 24	2.298 69
2.310 10	2.299 40
2.310 28	2.298 49
_	2.298 89
Average 2.310 10 ₉	2.299 47 ₂
Standard deviation 0.000 14 ₃	0.001 37 ₉

Example: Is Rayleigh's N_2 from Air Denser Than N_2 from Chemicals? (2 of 4)

Solution: The *F*-test told us that the standard deviations are significantly different, so we use Equations 4-9b and 4-10b:

$$t_{\text{calculated}} = \frac{\left|\overline{x}_{1} - \overline{x}_{2}\right|}{\sqrt{\left(s_{1}^{2}/n_{1}\right) + \left(s_{2}^{2}/n_{2}\right)}} = \frac{\left|2.310\ 10_{9} - 2.299\ 47_{2}\right|}{\sqrt{0.000\ 14_{3}^{2}/7} + 0.001\ 37_{9}^{2}/8} = 21.7$$

$$\text{Degrees of freedom} = \frac{\left(s_{1}^{2}/n_{1} + s_{2}^{2}/n_{2}\right)^{2}}{\left(\frac{s_{1}^{2}}/n_{1}\right)^{2}} + \frac{\left(s_{2}^{2}/n_{2}\right)^{2}}{n_{2} - 1}} = \frac{\left(0.000\ 14_{3}^{2}/7 + 0.001\ 37_{9}^{2}/8\right)^{2}}{\left(0.001\ 43_{9}^{2}/7\right)^{2}} + \frac{\left(0.001\ 37_{9}^{2}/8\right)^{2}}{8 - 1} = 7.17$$

Example: Is Rayleigh's N_2 from Air Denser Than N_2 from Chemicals? (3 of 4)

Solution: Equation 4-10b gives us 7.17 degrees of freedom, which we round to 7. For 7 degrees of freedom, the critical value of *t* in Table 4-4 for 95% confidence is 2.365. The observed value $t_{calculated} = 21.7$ far exceeds t_{table} . The obvious difference between the two data sets in Figure 4-7 is highly significant.



Figure 4-7

Example: Is Rayleigh's N_2 from Air Denser Than N_2 from Chemicals? (4 of 4)

Test Yourself: If the difference between the two mean values were half as great as Rayleigh found, but the standard deviations were unchanged, would the difference still be significant?

Case 3: Paired t Test for Comparing Individual Differences

Do the methods give the same answer?

- Use two methods to make single measurements on several different samples..
- No measurement is duplicated. $s_d = \sqrt{\frac{n-1}{n-1}}$

$$=\sqrt{\frac{(0.01-\overline{d})^2+(0.37-\overline{d})^2+(-0.14-\overline{d})^2+\cdots}{8-1}}=0.40$$

$t_{\text{calculated}} = \frac{|\vec{d}|}{s_d} \sqrt{n} = \frac{|0.11_4|}{0.40_1} \sqrt{8} = 0.80_3$

Figure 4-8

	A	В	С	D
1	Comparison of m	ethods for measuring ni	trate	
2				
3		Nitrate (ppm)	in plant extract	
4		Spectrophotometry	Experimental	
5	Sample	with Cd reduction	biosensor	Difference (d _i)
6	1	1.22	1.23	0.01
7	2	1.21	1.58	0.37
8	3	4.18	4.04	-0.14
9	4	3.96	4.92	0.96
10	5	1.18	0.96	-0.22
11	6	3.65	3.37	-0.28
12	7	4.36	4.48	0.12
13	8	1.61	1.70	0.09
14			Mean difference =	0.114
15		Standard de	eviation of differences =	0.401
16			Number of samples =	8
17	D6 = C6-B6		t _{calculated} =	0.803
18	D14 = AVERAGE	(D6:D13)	t _{table} =	2.365
19	D15 = STDEV.S(I	D6:D13)		
20	D16 = COUNT(D6	S:D13)		
21	D17 = ABS(D14)*	SQRT(A13)/D15 (ABS	= absolute value)	
22	D18 = T.INV.2T(0	.05,A13-1)		

One-Tailed and Two-Tailed Significance Tests

Two tailed tests: t test calculations assume:

• Certified value lies in the **outer 5%** of the area under the curve

One tailed tests: compare mean with regulatory limit

• 5% region lies only on one side of the certified mean

Consider drinking water: We are concerned only if the probability of arsenic (As) in water exceeds **the limit**. EPA maximum permissible level = 10 μ g As/L Water samples \rightarrow 10.06, 10.12, 10.19, and 10.04 μ g As/L \Rightarrow 10.10₂₅ \pm 0.06₇₅ μ g/L

$$t_{\text{calculated}} = \frac{|\bar{x} - \text{regulatory limit}|}{s} \sqrt{n} = \frac{|10.10_{25} - 10|}{0.06_{75}} \sqrt{4} = 3.0_{4}$$

Figure 4-9



Section 4-5 *t* Tests with a Spreadsheet

t Tests with a Spreadsheet

Spreadsheet for comparing mean values of Rayleigh's nitrogen measurements

Figure 4-

	A	В	С	D	E	F	G		
1	Analysis of Raylei	gh's Data			t-Test: Two-Sample Assum	t-Test: Two-Sample Assuming Equal Variances			
2						Variable 1	Variable 2		
3		Mass of gas (g) c	ollected from		Mean	2.310109	2.299473		
4		air	chemical		Variance	2.03E-08	1.9E-06		
5		2.31017	2.30143		Observations	7	8		
6		2.30986	2.29890		Pooled Variance	1.03E-06			
7		2.31010	2.29816		Hypothesized Mean Diff	0			
8		2.31001	2.30182		df	13			
9		2.31024	2.29869		t Stat	20.21372			
10		2.31010	2.29940		P(T<=t) one-tail	1.66E-11			
11		2.31028	2.29849		t Critical one-tail	1.770932			
12			2.29889		P(T<=t) two-tail 3.32				
13	Average	2.31011	2.29947		t Critical two-tail	cal two-tail 2.160368			
14	Std Dev	0.00014	0.00138						
15					t-Test: Two-Sample Assur	ning Unequal Varian	ces		
16	B13 = AVERAGE(B5:B12)				Variable 1	Variable 2		
17	B14 = STDEV.S(B	5:B12)			Mean	2.310109	2.299473		
18					Variance	2.03E-08	1.9E-06		
19					Observations	7	8		
20					Hypothesized Mean Diff	0			
21					df	7			
22					t Stat	21.68022			
23					P(T<=t) one-tail	5.6E-08			
24					t Critical one-tail	1.894578			
25					P(T<=t) two-tail	1.12E-07			
26					t Critical two-tail	2.364623			

Section 4-6 Grubbs Test for an Outlier

Grubbs Test: Check for Outliers

Should a data point that looks like an anomaly be discarded?

If you make several replicate measurements, results should fall within a Gaussian distribution about the mean. But when *n* is small, it can be difficult to determine if an outlying data point falls within the normal distribution.

The **Grubbs test** is a statistical test to decide whether to discard a datum that appears discrepant (an *"outlier"*).

If $G_{\text{calculated}} > G_{\text{table}}$, then reject the null hypothesis.

- There is <5% chance that the suspicious data point is a member of the same population as the other measurements.
- The difference is considered significant.

Grubbs Test for an Outlier (1 of 3)

When a single data point lies far from the other data in a set of measurements:

- First, check your notebook.
- Are there any recorded observations about the anomalous data point (for example, a note that solution was lost during transfer)?
- Any data point based on *recorded* faulty procedure should be discarded, no matter how well it fits the rest of the data (*"blunder"*).



Grubbs Test for an Outlier (2 of 3)

In the absence of a recorded blunder, use the Grubbs test.

$$G_{\text{calculated}} = \frac{|\text{questionable value} - \overline{x}|}{s}$$

$$G_{\text{calculated}} = \frac{|\text{questionable$$

Questionable value

Moon

- If $G_{calculated}$ is greater than G in Table 4-6, the questionable point should be discarded.
- Only one outlier may be rejected using the Grubbs test.

Grubbs Test for an Outlier (3 of 3)

In the absence of a recorded blunder, use the **Grubbs test**.

$$G_{\text{calculated}} = \frac{|\text{questionable value} - \overline{x}|}{s}$$

Volumes for replicate titrations (mL): 28.54, 28.39, 28.47, 27.68



$$\overline{\mathbf{x}} = \mathbf{28.27} \pm \mathbf{0.40} \text{ mL}; \quad \mathbf{RSD} = \mathbf{1.4\%} \leftarrow \begin{array}{l} \text{(Larger than expected} \\ \text{precision for this titration)} \end{array}$$

$$G_{\text{calculated}} = \frac{|27.68 - 28.27|}{0.40} = 1.482 \qquad G_{\text{table}} = 1.463 \text{ for 4 observations} \end{array}$$

 $G_{\text{calculated}} > G_{\text{table}}$, so reject the null hypothesis.

• The questionable data point is an "outlier" and should be discarded.

Table 4-6 Critical values of G for rejection of outlier

Number of observations	G (95% confidence)	Number of observations	G (95% confidence)
3	1.153	10	2.176
4	1.463	11	2.234
5	1.672	12	2.285
6	1.822	15	2.409
7	1.938	20	2.557
8	2.032	30	2.745
9	2.110	50	2.956

 $G_{calculated} = |questionable value - mean|/s.$ If $G_{calculated} > G_{table}$, the value in question can be rejected with 95% confidence. Values in this table are for a one-tailed test, as recommended by ASTM.

Section 4-7 The Method of Least Squares

The Method of Least Squares

For most chemical analyses, the response obtained by the given lab procedure must be compared to known quantities (called **standards**).

In this way the response from an *unknown* quantity can be interpreted.

- Prepare a calibration curve from known standards.
- Work in a region where the calibration curve is linear (usually).

Method of least squares: used to draw the "best" straight line through experimental data points that contain some scatter

- Some points will lie above and some below the line.
- Equation y = mx + b can be used to quantify the unknown from

Finding the Equation of the Line

Assume:

- Uncertainty in y values is much greater than uncertainty in x values $(s_v >> s_x)$.
- Uncertainties of all *y* values are similar.

Draw a line to minimize vertical deviations between points and line.



• Vertical deviation $d_i = y_i - y = y_i - (mx_i + b)$

Figure 4-11



Harris/Lucy, *Quantitative Chemical Analysis*, 10e, © 2020 W. H. Freeman and Company

Deviations can be positive or negative. To minimize magnitude, irrespective of sign, squade the deviation method of least squares

Determinants

Mathematically finding values of *m* and *b* that minimize the sum of the squares involves some calculus.

We express the final solution for *m* and *b* as **determinants**.



$$\begin{vmatrix} e & f \\ g & h \end{vmatrix} = eh - fg$$

$$\begin{vmatrix} 6 & 5 \\ 4 & 3 \end{vmatrix} = (6 \times 3) - (5 \times 4) = 2$$



Harris/Lucy, *Quantitative Chemical Analysis*, 10e, © 2020 W. H. Freeman and Company

Determinants to Solve Method of Least Squares



where D is

$$D = \begin{vmatrix} \sum (x_i^2) & \sum x_i \\ \sum x_i & n \end{vmatrix}$$

and *n* is the number of points.

Table 4-7 Calculations for least-squares analysis

x _i	y _i	$x_i y_i$	x_i^2	$d_i (= y_i - mx_i - b)$) d_i^2
1	2	2	1	0.038 46	0.001 479 3
3	3	9	9	$-0.192\ 31$	0.036 982
4	4	16	16	0.192 31	0.036 982
6	5	30	36	-0.03846	0.001 479 3
$\overline{\Sigma x_i} = 14$	$\overline{\Sigma y_i} = 14$	$\overline{\Sigma(x_i y_i)} = 57$	$\overline{\sum(x_i^2)} = 62$		$\sum (d_i^2) = 0.076923$

$$m = \begin{vmatrix} 57 & 14 \\ 14 & 4 \end{vmatrix} \div \begin{vmatrix} 62 & 14 \\ 14 & 4 \end{vmatrix} = \frac{(57 \times 4) - (14 \times 14)}{(62 \times 4) - (14 \times 14)} = \frac{32}{52} = 0.615 \ 38 \end{vmatrix}$$
$$b = \begin{vmatrix} 62 & 57 \\ 14 & 14 \end{vmatrix} \div \begin{vmatrix} 62 & 14 \\ 14 & 4 \end{vmatrix} = \frac{(62 \times 4) - (57 \times 14)}{(62 \times 4) - (14 \times 14)} = \frac{70}{52} = 1.346 \ 15 \end{vmatrix}$$

y = 0. 615 38 *x* + 1. 346 15

Example: Finding Slope and Intercept with a Spreadsheet (1 of 3)

Excel has functions called SLOPE and INTERCEPT, whose use is illustrated here:

	Α	В	С	D	E	F	
1	х	у			Formulas:		
2	1	2		slope =			
3	3	3		0.61538 D3 = SLOPE(B2:B5,A2:A5)			
4	4	4		intercept =			
5	6	5		1.34615 D5 = INTERCEPT(B2:B5,A2:A5)			

Example: Finding Slope and Intercept with a Spreadsheet (2 of 3)

The slope in cell D3 is computed with the formula "=SLOPE(B2:B5,A2:A5)", where B2:B5 is the range containing the y values and A2:A5 is the range containing x values.

	Α	В	С	D	E	F	
1	х	у			Formulas:		
2	1	2		slope =			
3	3	3		0.61538	0.61538 D3 = SLOPE(B2:B5,A2:A5)		
4	4	4		intercept =			
5	6	5		1.34615 D5 = INTERCEPT(B2:B5,A2:A5)			

Example: Finding Slope and Intercept with a Spreadsheet (3 of 3)

Test Yourself: Change cell A3 from 3 to 3.5 and find the new slope and intercept.

How Reliable Are Least Squares Parameters?

To estimate uncertainties in *m* and *b*, an uncertainty analysis must be performed.

Estimate σ_y , the population standard deviation for all y, by calculating s_y .

$$s_{y} = \sqrt{\frac{\sum(d_{i}^{2})}{n-2}} \qquad u_{m}^{2} = \frac{s_{y}^{2}n}{D} \qquad u_{b}^{2} = \frac{s_{y}^{2}\sum(x_{i}^{2})}{D}$$
The first digit of the uncertainty is the last significant figure. We often retain extra, insignificant, digits to prevent round-off errors in further calculations. Intercept:
$$\begin{array}{c} 1.346\ 15\\ \pm 0.214\ 15 \end{array} = 0.62\ \pm 0.05\ or\ 0.61_{5}\ \pm 0.05_{4} \end{array}$$

Example: Finding s_y , u_m , and u_b with a Spreadsheet (1 of 3)

Excel function LINEST returns the slope and intercept and their standard uncertainties in a table (a *matrix*). As an example, enter x and y values from Table 4-7 in columns A and B. Then highlight the 3-row × 2-column region E3:F5 with your mouse. This block of cells is selected for the output of LINEST. On the Formulas ribbon, go to Insert Function. In the window that appears, in "Or select a category" select Statistical and double click LINEST. The new window asks for four inputs to the function. For y values, enter B2:B5. Then enter A2:A5 for x values. The next two entries are both "TRUE". The first TRUE tells Excel that we want to compute the y-intercept of the line and not force the intercept to be 0. The second TRUE tells Excel to print out the uncertainties as well as the slope and intercept. The formula you have just entered is "=LINEST(B2:B5,A2:A5,TRUE,TRUE)". Now press CONTROL+SHIFT+ENTER on a PC or CONTROL+SHIFT+RETURN on a Mac.

Example: Finding s_y , u_m , and u_b with a Spreadsheet (2 of 3)

Excel prints out a matrix in cells E3:F5. Write labels around the block to indicate what is in each cell. The slope and intercept are on the top line. The second line contains u_m and u_b . Cell F5 contains s_y , and cell E5 contains a quantity called R^2 , which is defined in Equation 5-3 and is a measure of the goodness of fit of the data to the line. The closer R^2 is to unity, the better the fit.

	Α	В	С	D	E	F	G
1	х	у			Output fro	m LINEST	
2	1	2			Slope	Intercept	
3	3	3		Parameter	0.61538	1.34615	
4	4	4		u _m	0.05439	0.21414	u _b
5	6	5		R ²	0.98462	0.19612	s _v
6	High	light	cells	E3:F5			
7	Туре	JE)"					
8	Pres						
9	Pres	s CTI	RL+S	HIFT+RETUR	RN (on Mac)		

Example: Finding s_y , u_m , and u_b with a Spreadsheet (3 of 3)

Test Yourself: Change cell A3 from 3 to 3.5 and apply LINEST. What is the value of s_v from LINEST?

Section 4-8 Calibration Curves

Calibration Curves

A **calibration curve** shows the response of an analytical method to known quantities of analyte.

- **Standard solutions** contain known concentrations of analyte.
- **Blank solutions** contain all reagents and solvents used in the analysis, but contain *no* deliberately added analyte.

A spectrophotometer measures the absorbance of light (*y*-axis), which is proportional to the quantity of protein analyzed (*x*-axis).



Table 4-8 Spectrophotometer data used toconstruct calibration curve

Amount of protein	4	Absorbance o	of	Dente			
(µg)	Indep	pendent stan	dards	Range	Corr	ected absorb	ance
0	0.099	0.099	0.100	0.001	-0.000 ₃	-0.000 ₃	0.000 ₇
5.0	0.185	0.187	0.188	0.003	0.085 ₇	0.087 ₇	0.088 ₇
10.0	0.282	0.272	0.272	0.010	0.182 ₇	0.172 ₇	0.172 ₇
15.0	0.345	0.347	(0.392)	0.047	0.245 ₇	0.247 ₇	_
20.0	0.425	0.425	0.430	0.005	0.325 ₇	0.325 ₇	0.330 ₇
25.0	0.483	0.488	0.496	0.013	0.383 ₇	0.388 ₇	0.396 ₇

Constructing a Calibration Curve (1 of 2)

- 1. Prepare known samples of analyte covering the range (0 to 150%) of concentrations expected for unknowns.
 - Tabulate amount of analyte in each standard and response.
- 2. Subtract the average absorbance of the blank solutions from each measured absorbance (corrected absorbance).
 - Blanks measure the response of the procedure when no analyte is present.
- 3. Make a graph of corrected absorbance vs. quantity of analyte.
 - Inspect the graph for linearity, outliers, and consistent *y*-uncertainty.

Figure 4-12



Constructing a Calibration Curve (2 of 2)

- 4. Use the least-squares procedure to find the best straight line through the linear portion of the data. Corrected absorbance = $(0.016 \ 3_0)(\mu g \text{ protein}) + 0.004_7$
- 5. If you analyze an unknown at a future time, run a blank *at that time*.
 - Subtract the new blank signal from the unknown to correct.

Figure 4-12



Example: Using a Linear Calibration Curve (1 of 3)

An unknown protein sample gave an absorbance of 0.406, and a blank had an absorbance of 0.104. How many micrograms of protein are in the unknown?

$$\underbrace{\text{Corrected absorbance}}_{y} = (0.0163_0)\underbrace{(\mu g \text{ protein })}_{x} + 0.004_7$$

Example: Using a Linear Calibration Curve (2 of 3)

Solution: The corrected absorbance is 0.406 – 0.104 = 0.302, which lies on the linear portion of the calibration curve in Figure 4-13. Rearranging Equation 4-25 gives:

 $\mu g \text{ of protein} = \frac{\text{corrected absorbance } -0.004_7}{0.0163_0}$ $= \frac{0.302 - 0.004_7}{0.0163_0} = 18.2_4 \ \mu g$

Figure 4-13



Example: Using a Linear Calibration Curve (3 of 3)

Test Yourself: What mass of protein gives a corrected absorbance of 0.250?

<u>Corrected absorbance</u> = $(0.0163_0)(\mu g \text{ protein}) + 0.004_7$

Linear Response

The **linear range** of an analytical method is the analyte concentration range over which response is proportional to concentration.

Dynamic range is the concentration range over which there is a measurable response to analyte, even if the response is not linear.

- Calibration procedures with a *linear response* are preferred.
- Corrected analytical signal ∞ quantity of analyte.
- It is possible to obtain valid results beyond the linear region by fitting with a nonlinear equation.

Figure 4-14



Box 4-2 Using a Nonlinear Calibration Curve

- Consider an unknown whose corrected absorbance of 0.375 lies beyond the linear range.
- Fit all the data points with a quadratic equation:

 $y = -1.1_7 \times 10^{-4} x^2 + 0.01855_8 x - 0.0007$

• Insert *y* = 0.375 into the equation and rearrange to the form

$$ax^2 + bx + c = 0$$

• Solve for *x*.

$$x = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a} = 135 \,\mu \text{g or } 23_{.8} \,\mu \text{g}$$





Harris/Lucy, *Quantitative Chemical Analysis*, 10e, © 2020 W. H. Freeman and Company

Good Practice (1 of 2)

Always make a graph of your data



- Helps reject bad data, stimulus to repeat a measurement, or decision that a straight line is not appropriate
- All three data sets were fit to y = 0.5x + 3

Good Practice (2 of 2)

- It is not reliable to extrapolate any calibration curve beyond the measured range of standards.
- At *least* six calibration concentrations and two replicate measurements of each unknown are recommended.
- Make each standard solution from a certified material.
- Avoid serial dilution of a single stock solution (serial dilution propagates systematic error).
- Measure calibration solutions in random order, not in consecutive order of increasing concentration.

Box 4-3 Importance of Graphs to Visualize Data

A good graph reveals key characteristics of data and guides statistical analysis.



- Heights on the bar graph (a) give the mean values of two data sets.
 The error bars correspond to ±standard deviation of the mean.
- Data plots (b–e) show different characteristics of the data that are not evident in the bar graph.

Propagation of Uncertainty with a Calibration Curve

- An unknown with a corrected absorbance of y = 0.302 had a protein content of $x = 18.2_4 \mu g$. What is the uncertainty in x?
- Standard uncertainty in *x* = standard deviation of the mean =

$$u_{x} = \frac{s_{y}}{|m|} \sqrt{\frac{1}{k} + \frac{1}{n} + \frac{(y - \overline{y})^{2}}{m^{2} \sum (x_{i} - \overline{x})^{2}}}$$
$$u_{y} = \pm 0.2_{3} \,\mu g$$

k = number of replicate measurementsn = number of data points

Confidence interval for x is ±tu_x, where t is Student's t for n - 2 degrees of freedom

$$\pm tu_x = \pm (2.179)(0.2_3) = \pm 0.5_0 \,\mu g$$

Section 4-9 A Spreadsheet for Least Squares

Figure 4-16: Spreadsheet for Linear Least-Squares Analysis

	A	В	С	D	E		F	G	Н	I.
1	Least-Squares Spreadsheet									
2					[6				
3	Highlight cells B10:C12	х	у			0				
4	Type "= LINEST(C4:C7,	1	2					y = 0.6154x +	1.3462	
5	B4:B7,TRUE,TRUE)	3	3			5				
6	For PC, press	4	4							
7	CTRL+SHIFT+ENTER	6	5			4				
8	For Mac, press									
9	COMMAND+RETURN	LINEST ou	utput:							
10	m	0.6154	1.3462	b		> 3				
11	u _m	0.0544	0.2141	u _b						
12	R ²	0.9846	0.1961	s _y		2				+ I C
13										
14	n =	4	B14 = COUNT(B4:B7)			1				
15	Mean y =	3.5	B15 = AVE	ERAGE(C4:C	7)					
16	$\Sigma(x_i - \text{mean } x)^2 =$	13	B16 = DE	/SQ(B4:B7)						
17						0	D 1	2 3	3 4	5 6
18	Measured y =	2.72	Input)	c	
10	k = Number of replicate				L					_
19	measurements of y =	1	Input							
20	Derived x =	2.2325	B20 = (B1	8-C10)/B10						
21	u _x =	0.3735	B21 = (C1)	2/ABS(B10))*	SQRT((1/E	319)+(1/B14)+((B18-B15)^2)	/(B10^2*B16))

Figure 4-17: Adding Error Bars to a Graph

